



Data setup for linear models

We assume $Y \sim N(\mu, \sigma^2)$ and the data are independent.

 μ and σ^2 are parameters that describe centre and spread of the distribution respectively.

We can shift the normal distribution without changing its shape.

This allows us to split the data into 2 pieces, one fixed, the other random.

- $Y_i = \mu + \varepsilon_i$
- μ is a fixed value and ε is the random error in our data.

We assume $\varepsilon \sim N(0, \sigma^2)$

We can model both the fixed and random portions of our data.

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Estimating Fixed Effects

We do not know the true value of the fixed effects but we can estimate it using the data.

The estimates of the parameters will follow a normal distribution because of the Central Limit Theorem.

 $\hat{b} \sim N(b, se_{\hat{b}}^2)$

The standard error (se) associated with \hat{b} is related to σ , the sample size, the spread of x_{ij} and its correlation with the other predictors.

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There is a 95% chance that the true value of b is between $\hat{b}\pm 1.96 se_{\hat{h}}.$



Random Effects continued

Because we selected schools, classes and students randomly from a larger group we treat these effects as random. Since these effects are nested, they form a hierarchical linear model.

The random effects assumed to be independent with

 $\gamma \sim N(0, \sigma_s^2)$ (Between school variation). $\delta \sim N(0, \sigma_c^2)$ (Between class but within school variation). $\varepsilon \sim N(0, \sigma^2)$ (Between student but within class variation).

This model allows us to estimate how much variation in aptitude score between students is due to the school they attend and the class they are in.

Random Effects can depend on parameters as well. Anything beyond a random intercept model is a big jump in complexity.

Random Effects continued

Random effects models assume data are conditionally independent given the random effects.

Unconditionally, observations on the same unit are correlated.

If we have $Y_{ij} = \mu + \gamma_i + \delta_{ij}$ being the *j*th observation on the *i*th subject.

 $\gamma \sim \textit{N}(0,\sigma_B^2)$ and $\delta \sim \textit{N}(0,\sigma_W^2)$

We can compute the correlation between observations on the same and different subjects.

$$Cor(Y_{1,1}, Y_{2,1}) = Cor((\gamma_1 + \delta_{11}), (\gamma_2 + \delta_{21})) = 0$$

$$Cor(Y_{1,1}, Y_{1,2}) = Cor((\gamma_1 + \delta_{11}), (\gamma_1 + \delta_{12})) = \frac{\sigma_B^2}{\sigma_B^2 + \sigma_W^2}$$

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Fixed Effects model for within subject factor

Subject is included as a fixed blocking factor.

Anova Table, Response = distance ## Df Sum Sq Mean Sq F value Pr(>F) ## Age 3 237.19 79.064 38.0396 2.986e-15 ## Subject 26 518.38 19.938 9.5925 3.375e-15 ## Residuals 78 162.12 2.078 ## Estimate Std. Error t value ## (Intercept) 21.162 0.760 27.85 ## Age10 2.50 0.981 0.392 ## Age12 2.463 0.392 6.28 ## Age14 3.907 0.392 9.96 ・ロト・日本・ キャー キー うくぐ

Mixed Effects model for within subject factor								
Subject is included as a random effect.								
## Anova Table, Response = distance								
## NumDF DenDF F.value Pr(>F)								
## Age 3 78 38.04 2.998e-15								
-								
## Var SD								
## Subject 4.465 2.113								
## Residual 2.078 1.442								
## Estimate Std. Error	df t value							
## (Intercept) 22.185 0.492 43	3.4 45.07							
## Age10 0.981 0.392 78	8.0 2.50							
## Age12 2.463 0.392 78	8.0 6.28							
## Age14 3.907 0.392 78	8.0 9.96							
-								





Mixed Effects model approach

Mixed effects model works without having to average and is accurate even if there is imbalance over the subjects.

```
## Anova Table, Response = distance
##
       NumDF DenDF F.value
                            Pr(>F)
## Sex
          1
               25 9.2921 0.005375
##
              Var
                     SD
## Subject 2.547 1.596
## Residual 4.930 2.220
              Estimate Std. Error df t value
##
                 24.97
                            0.486 25
                                       51.38
## (Intercept)
## SexFemale
                 -2.32
                            0.761 25
                                       -3.05
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```







Mixed Effects Model approach

Anova Table, Response = distance NumDF DenDF F.value ## Pr(>F) 78 38.040 2.998e-15 ## Age 3 ## Sex 25 9.292 0.005375 1 ## SD Var ## Subject 3.260 1.805 ## Residual 2.078 1.442 ## Estimate Std. Error df t value 23.131 0.542 38 42.66 ## (Intercept) 2.50 ## Age10 0.981 0.392 78 ## Age12 2.463 0.392 78 6.28 ## Age14 3.907 0.392 78 9.96 ## SexFemale -2.321 0.761 25 -3.05 ・ロト・日本・山田・ 山田・ 山田・



```
Within subject factor (Unbalanced case)
   Fixed Effects Model
   ## Anova Table, Response = distance
   ##
                Df Sum Sq Mean Sq F value
                                             Pr(>F)
                 3 223.94 74.646 48.577 4.676e-16
   ## Age
   ## Subject
                26 414.28 15.934 10.369 6.920e-14
   ## Residuals 60 92.20 1.537
   Mixed Effects Model
   ## Anova Table, Response = distance
          NumDF DenDF F.value
   ##
                                  Pr(>F)
   ## Age
              3 60.027 36.312 1.643e-13
   Either model is fine but the choice affects the interpretation.
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Between subject factors (Unbalanced case)
Fixed Effects model on average observation
<pre>## Anova Table, Response = average distance ## Df Sum Sq Mean Sq F value Pr(>F) ## Sex 1 57.085 57.085 12.607 0.001554 ## Residuals 25 113.196 4.528</pre>
Mixed Effects Model on raw data
<pre>## Anova Table, Response = distance ## NumDF DenDF F.value Pr(>F) ## Sex 1 22.776 13.035 0.001489</pre>
Mixed effects takes variation in sample size into account. Result is more accurate.
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Together in a Mixed Effects model

Anova Table, Response = distance ## NumDF DenDF F.value Pr(>F) 3 59.947 35.703 2.305e-13 ## Age ## Sex 1 23.960 10.980 0.002918 ## SD Var ## Subject 3.478 1.865 ## Residual 1.553 1.246 ## Estimate Std. Error df t value 23.37 42.26 ## (Intercept) 0.553 35.4 ## Age10 1.22 0.408 60.7 3.00 ## Age12 2.35 0.369 59.8 6.36 ## Age14 3.79 0.382 59.8 9.92 ## SexFemale -2.590.782 24.0 -3.31 ・ロト・日本・山田・ 山田・ 山田・





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The A	Anova table	2					
## ## ## ## ## ##	Anova Table Nu Sex Minority SES Size Sector PRACAD DISCLIM	e, Re imDF 1 1 1 1 1 1 1	esponse DenDF 5148.8 2734.6 6776.5 154.0 148.1 165.1 160.3	= Math // F.value 62.36 240.77 348.18 14.31 4.08 39.87 4.06	Achievement Pr(>F) 3.553e-15 < 2.2e-16 < 2.2e-16 0.000221 0.045275 2.413e-09 0.045602	Score	
Nc val da DF siz	ote the denomi ued. This is go ta. ⁻ is a paramete e.	nator enera r in a	degrees Ily true n statistic	of freedon nixed effec al distribut	n (DF) are no ts models wit tion. It is relat	t integer h unbalanced ted to sample	~ ~ ~

The fixed and random effects

##	Var	SD					
##	School 1.462	2 1.209					
##	Residual 35.897	5.991					
##		Estimate	Std.	Error	df	t value	
##	(Intercept)	10.653		0.460	180	23.17	
##	SexFemale	-1.253		0.159	5149	-7.90	
##	MinorityYes	-3.043		0.196	2735	-15.52	
##	SES	1.971		0.106	6777	18.66	
##	Size	0.823		0.218	154	3.78	
##	SectorCatholic	0.802		0.397	148	2.02	
##	PRACAD	4.283		0.678	165	6.31	
##	DISCLIM	-0.375		0.186	160	-2.01	
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More than one Random Effect (Hierarchical or Crossed models)

The models above demonstrate that we can fit 2 levels of data in the same model using Mixed Effects model.

We can fit multiple levels simultaneously using this technique. Each level is defined by a random effect.

A hierarchical model requires the random effects be nested to create a hierarchy.

 $Y_{ijk} = \mu + \gamma_i + \delta_{ij} + \varepsilon_{ijk}$

However random effects do not have to be nested they can be crossed.

$$Y_{ijk} = \mu + \gamma_i + \delta_j + \varepsilon_{ijk}$$

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The fi	The fixed effects						
The	The estimated group means are below:						
##		service	dept	Estimate	SE	DF	
##	1	FALSE	6	54.7	1.19	502	
##	2	TRUE	6	50.2	1.17	453	
##	3	FALSE	4	57.1	1.07	560	
##	4	TRUE	4	56.4	1.14	700	
##	5	FALSE	9	55.9	1.45	376	
##	6	TRUE	9	50.0	1.55	465	
The	e fi	ixed effects	s show	that the le	cture r	ratings are similar for Dept 6	
and 9 but are higher for Dept 4. In general service courses receive a							
lower rating within the 3 departments.							
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he random effects and correlations								
The variance components of the model are below:								
## Var SD ## Student 49.80 7.057 ## Instructor 98.35 9.917 ## Residual 584.59 24.178								
Observations from difference students on different instructors are uncorrelated.								
Same student on different instructors: $ ho=0.08$								
Different students on same instructor: $ ho=0.14$								
Each instructor/student combination appeared only once in the dataset so there is no same student same instructor correlation estimate.								
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Mixed Effects model for non-normal responses.

For other distributions, the mean and variance are related and cannot be separated

The link function relates the fixed effects to the mean and transforms the mean to the entire real line. We also put the random effects within the link function.

Dist	Link	Variance	Dist	Link	Variance
Poisson	$\log(\mu)$	μ	NB	$\log(\mu)$	$\mu + \mu^2/r$
Binomial	$logit(\mu)$	$\mu(1-\mu)$	NB	$\log(\mu)$	$\tau^2 \mu$

$$g(\mu_{ij}) = \alpha + \beta x_{ij} + \gamma_i$$
 where $\gamma_i \sim N(0, \sigma^2)$.

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Example: Contagious bovine pleuropneumonia (CBPP) We have repeated measures on herd of zebu cattle in Ethiopia. There are 15 herds with 3 or 4 repeated measures. The response is the number of cattle with CBPP.							
##	Likelihood H	Ratio Test					
## ##	period 3 25	5.61 1.151e-05					
	1						
##	Var	SD					
##	herd 0.4123	0.6421					
##		Estimate Std.	Error	z value	Pr(> z)		
##	(Intercept)	-1.398	0.231	-6.05	1.47e-09		
##	period2	-0.992	0.303	-3.27	1.07e-03		
##	period3	-1.128	0.323	-3.49	4.74e-04		
##	period4	-1.580	0.422	-3.74	1.82e-04		
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Questions?

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